Genetic Modulation of GABA Levels in the Anterior Cingulate Cortex by *GAD1* and *COMT*

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γ-Aminobutyric acid (GABA)-ergic transmission is critical for normal cortical function and is likely abnormal in a variety of neuropsychiatric disorders. We tested the *in vivo* effects of variations in two genes implicated in GABA function on GABA concentrations in prefrontal cortex of living subjects: glutamic acid decarboxylase 1 (*GAD1*), which encodes GAD67, and catechol-*o*-methyltransferase (*COMT*), which regulates synaptic dopamine in the cortex. We studied six single nucleotide polymorphisms (SNPs) in *GAD1* previously associated with risk for schizophrenia or cognitive dysfunction and the val158met polymorphism in *COMT* in 116 healthy volunteers using proton magnetic resonance spectroscopy. Two of the *GAD1* SNPs (rs1978340 ($p = 0.005$) and rs769390 ($p = 0.004$)) showed effects on GABA levels as did *COMT* val158met ($p = 0.04$). We then tested three SNPs in GAD1 (rs1978340, rs11542313, and rs769390) for interaction with COMT val158met based on previous clinical results. In this model, rs11542313 and *COMT* val158met showed significant main effects ($p = 0.001$ and 0.003, respectively) and a trend toward a significant interaction ($p = 0.05$). Interestingly, *GAD1* risk alleles for schizophrenia were associated with higher GABA/Cre, and Val-Val homozygotes had high GABA/Cre levels when on a *GAD1* risk genotype background ($N = 6$). These results support the importance of genetic variation in *GAD1* and *COMT* in regulating prefrontal cortical GABA function. The directionality of the effects, however, is inconsistent with earlier evidence of decreased GABA activity in schizophrenia.

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INTRODUCTION

γ-Aminobutyric acid (GABA)-ergic function is critical for numerous aspects of cognition (Gao and Goldman-Rakic, 2003a; Goldman-Rakic, 1990; Menzies *et al*, 2007; Seamans and Yang, 2004), and its alterations have been implicated in a number of neuropsychiatric disorders, including epilepsy (Wong *et al*, 2003), schizophrenia (Lewis and Hashimoto, 2007; Lewis *et al*, 2005), and depression (Sanacora *et al*, 1999). Studies of GABA activity *in vivo* in humans, however, have been very limited and consequently characterizing the basic physiology of the system *in vivo* has been an elusive goal.

Genes that contribute to the metabolism of GABA and the activity of GABA neurons are critical in GABA function, as they affect neurodevelopment and abnormalities in their function are likely to predispose to neuropsychiatric disorders. One of the synthetic enzymes for GABA, glutamic acid decarboxylase 67 (GAD67, encoded by *GAD1* on 2q31, a 46 kb gene consisting of 17 exons), is particularly relevant in this context because single nucleotide polymorphisms (SNPs) in this gene have been associated with GAD1 mRNA levels (Mellios *et al*, 2009; Straub *et al*, 2007), with risk for schizophrenia (Addington *et al*, 2005; Du *et al*, 2008; Straub *et al*, 2007; Zhao *et al*, 2007), with individual measures of neuroticism (Hettema *et al*, 2006) as well as with measures of prefrontal cortical function in normal subjects (Straub *et al*, 2007). Moreover, decreases in the expression of this gene have been found in several post-mortem studies of schizophrenia (eg, Hashimoto *et al*, 2008; Huang and Akbarian, 2007a; Straub *et al*, 2007, reviewed in Akbarian and Huang, 2006). *GAD1* expression and related GABA levels have also been found to have a specific role in sculpting interneuron axon growth and synapse formation during development (Chattopadhyaya *et al*, 2007). Another gene with potential implications for GABA function and risk for schizophrenia is catechol-*o*-methyltransferase (*COMT*), an enzyme that regulates cortical DA levels (Gogos *et al*, 1998; Karoum *et al*, 1994) and likely GABA neuronal

excitability (Seamans *et al*, 2001), perhaps particularly during development (Tseng *et al*, 2007), and would thus be expected to modulate GABA function.

In this study, we tested whether several SNPs in *GAD1* previously associated with risk for schizophrenia alter GABA levels in the anterior cingulate cortex (ACC) measured *in vivo*. We further investigated the effects of the functional variant val158met in *COMT*. The study by Straub *et al* (2007) found evidence of statistical epistasis between *GAD1* and *COMT* in increasing risk for schizophrenia. We thus tested whether this specific interaction would be recapitulated in terms of GABA levels in healthy individuals.

Given the neuropathology and genetic studies of schizophrenia mentioned above, we hypothesized that GABA levels would be reduced in carriers of *GAD1* alleles that confer risk for schizophrenia and that there would be a significant interaction with *COMT*, where the combination of Val alleles (relatively disadvantageous to prefrontal functioning: Blasi *et al*, 2005; Egan *et al*, 2001; Tan *et al*, 2007; Winterer *et al*, 2006) and risk alleles in *GAD1* would lead to further reductions in GABA. This is the first study that addresses the impact of genetic variation in any enzyme on *in vivo* GABA measurements.

MATERIALS AND METHODS

Participants

We recruited 116 adult healthy volunteers (59 men, 57 women) who had participated in the 'CBDB sibling study of the genetics of schizophrenia' (NCT00001486). Subjects were all Caucasian (mean age = 32.8 ± 10.2 years, mean $\pm$ SD, range 18–55). Exclusion criteria included any current or history of psychiatric (DSM-IV axis I or II by a modified Structured Clinical Interview) or medical illness affecting the brain, pregnancy, and current psychotropic medication use. All subjects had no first- or second-degree relatives with a psychotic disorder according to an assessment performed with the Family Interview for Genetic Studies (http://zork.wustl.edu/nimh/digs/figs/). Participants were informed of the purpose of the study and written consent (NIMH protocols 95-M-0150 and 91-M-0124) was obtained. This selected population was chosen to avoid potential concomitant effects of medications, illness duration, and phase of illness that are problematic in any patient-based research. Only three subjects were smokers.

Neuroimaging

Participants were scanned on three 3 T scanners (GE, Waukesha, Wisconsin), two of which were equipped with a 14m4 software platform ($n = 79$ and 22), whereas one ran on LX software ($n = 15$). A quadrature transmit–receive head coil was used on all scanners (IGC-Medical Advances, Milwaukee, WI). A T1-weighted 3D SPGR sequence (TR/TE, 9.2/3.3 ms, flip angle 17°, FOV 24 mm, matrix 256 × 192, 76 slices, 2 mm thick) was used for voxel placement and for image segmentation with SPM5 (http://www.fil.ion.ucl.ac.uk/spm). This segmentation was used to obtain relative proportions of gray, white matter, and cerebrospinal fluid (CSF) present in the voxel by an automated program written in IDL (ITT Visual Information Solutions, Boulder, CO, USA) by ASB.

For magnetic resonance spectroscopy (MRS), a single voxel of interest ($2 \times 2 \times 4.5$ cm: 18 cm^3) was placed in the medial prefrontal cortex (PFC) to maximize coverage of the gray matter in the dorsal ACC (Figure 1). We chose this location because of the ease of measurement (a large amount of gray relative to white matter can be investigated with excellent shimming despite the high magnetic field) and because this region has been shown to function abnormally in schizophrenia, with several approaches (eg, (Adams and David, 2007; Camchong *et al*, 2008; Carter *et al*, 2001; Deicken *et al*, 1997; Ende *et al*, 2000; Ford *et al*, 2004; Fornito *et al*, 2008; Ko *et al*, 2009; Manoach *et al*, 2007; Segall *et al*, 2009), including post-mortem studies of GAD67 mRNA transcript (Woo *et al*, 2004). Moreover, this region has been shown to be important in conflict and error detection and conflict monitoring (eg, Botvinick *et al*, 1999; Carter *et al*, 1999) and GABA circuitry is likely to be key to its proper functioning. It might have been preferable to study the dorsolateral prefrontal cortex (DLPFC), rather than the medial PFC, however several technical difficulties discouraged us from doing so. Measurement of the DLPFC with the same signal to noise resulted in much broader line widths in pilot studies and ended up including about 20% of gray matter in the voxel *vs* the current 75% for the ACC.

An interleaved PRESS-based J-editing method (Hasler *et al*, 2007; Sailasuta *et al*, 2001) was used for measurement of metabolites (TR/TE 1500/68 ms, NEX = 2, 768 averages, 2048 acquisition points, sampling frequency = 5000 Hz, 20 min scanning time). This method is perhaps more sensitive to the larger intracellular pool of GABA produced by GAD67 (Asada *et al*, 1997; Tian *et al*, 1999) as *GAD2*, which encodes for GAD65, likely produces a much smaller
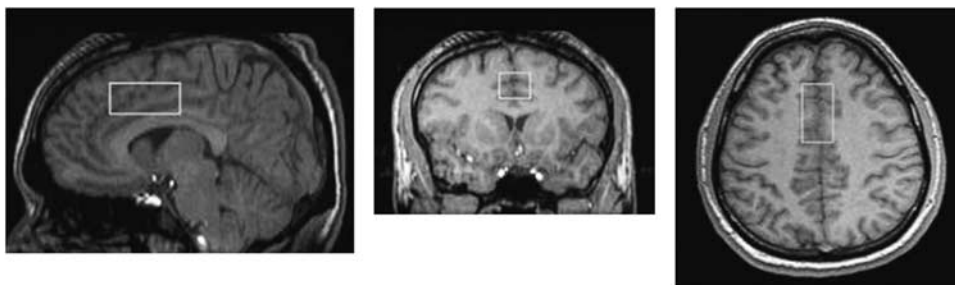

Figure 1 Magnetic resonance spectroscopy (MRS) voxel position. Sagittal, coronal, and axial images showing the position of the anterior cingulate voxel of interest, placed in the medial prefrontal cortex while trying to maximize gray matter.
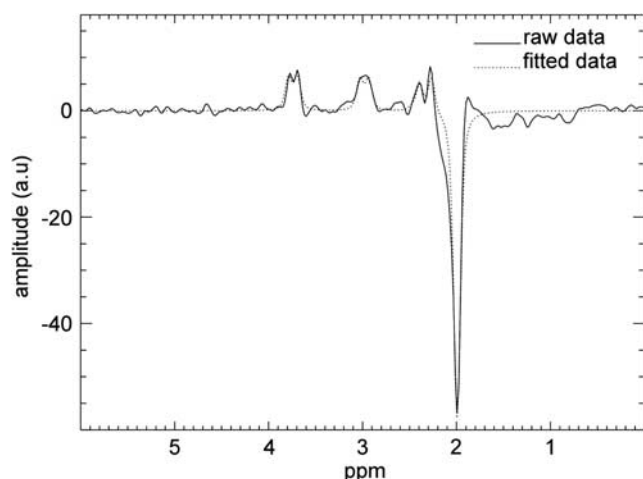
Figure 2 Spectroscopy output. A typical difference spectrum is displayed, showing the GABA resonance at 3.0 p.p.m. and co-edited Glx at 3.8 p.p.m. The *N*-acetyl-aspartate (NAA) signal at 2 p.p.m. appears inverted due to the editing that suppresses frequencies around this area of the spectrum. The subtraction of the unedited from the edited spectrum will thus result in a negative peak for NAA. The dotted line represents the automated fit to the raw data (continuous line).

pool of GABA used for vesicular release (review in Soghomonian and Martin, 1998). A fully automated non-linear fitting program written in IDL by JWvdV was used to calculate metabolite levels. The amplitudes of residual water, *N*-acetyl aspartate (NAA), choline (Cho), and creatine (Cre) were fitted in the unedited spectrum. Subsequently, the unedited spectrum was subtracted from the edited spectrum, resulting in a difference spectrum where the co-edited Glx (glutamate + glutamine) and GABA signals were separated from overlapping resonances. After subtraction of the water baseline, GABA and Glx were fitted automatically, using the linewidth and positions fit results from the fit of Cho, Cre, and NAA in the unedited spectrum. The entire process yielded estimates for absolute signal intensities and ratios of the metabolites to Cre and to water (see Figure 2 for an example of the output). The fitting procedure also rejected pairs of acquisitions where the differences in the absolute value of the residual water signals of a pair were more than 10%, indicating possible movement during the scan (van der Veen *et al*, 2007). Two further quality control procedures were put in place: (1) at the time of scanning, the voxel was shimmed to a linewidth for water below 10 Hz in all cases (the average linewidth being 7 ± 0.96 Hz); (2) our expert physicist (JWvdV) reviewed all spectra masked to genotype to ensure adequate fitting and line shapes. A separate estimate of Cre was obtained by processing the unedited portion of the spectrum with LCmodel (Provencher, 2001), a commonly used technique to obtain absolute quantitative estimates (derived using water as a reference). Cre values were further corrected for the amount of CSF in the voxel by dividing them by the proportion of tissue in the voxel (percent gray + percent white matter).

Genotyping

Ten SNPs were genotyped in *GAD1* (rs10432420, rs7557793, rs1978340, rs872123, rs3762555, rs3749034, rs2241165,

rs11542313 (formerly known as rs16823181), rs769390, and rs3791850) based on previous associations of these SNPs with risk for schizophrenia, *GAD1* mRNA expression, gray matter loss, and/or cognitive performance (Addington *et al*, 2005; Straub *et al*, 2007; see Table 1), rs7557793, rs872123, rs3762555, and rs2241165 were in strong linkage disequilibrium (LD) with rs3749034 (all $r^2 > 0.8$), so they were dropped from further analysis, leaving six SNPs to investigate further. rs3749034 was kept for analysis because previous evidence from our post-mortem sample indicated that this SNP was associated with GAD1 mRNA expression (Straub *et al*, 2007). Comparing single SNP and haplotype results with both clinical and cognitive phenotypes, we have found that when SNPs are in such strong LD (here all r^2 values are > 0.80) and the genotyping call rate is high, with our sample size, there is usually little advantage to including the highly redundant genotypes. Haplotype analysis would have added to the number of tests, and interpretation of results from rare haplotypes is difficult. Detailed information about the SNPs is presented in Table 1 and additional information regarding the LD structure is shown in Figure 3. Of the remaining SNPs, rs1978340, rs11542313, and rs769390 had been shown to increase risk for schizophrenia when the risk allele was present on a *COMT* Val homozygote background (Straub *et al*, 2007; Table 1). Thus, we explored specifically the potential interaction of these SNPs with COMT on GABA levels. *COMT* Val158Met (rs4680) was also genotyped. All genotyping was performed using the Taq-Man allelic discrimination assay, as previously described (Straub *et al*, 2007). All genotypes were in Hardy–Weinberg equilibrium. Genotype reproducibility was assessed by comparing genotypes obtained from plates used for this study to plates genotyped for Straub *et al* (2007). Depending on the SNP in question, we found that somewhere between 153 and 236 individuals had duplicate genotypes. Discordance rate was below 2.3%, and discordant genotypes were removed before analysis.

Statistics

Before assessing the effects of genotype on GABA/Cre (our main dependent variable), we addressed the effects of age, sex, scanner, and percent gray matter contribution to the voxel in the whole cohort. This was carried out by testing ANCOVA models with GABA/Cre as the dependent variable, scanner and gender as factors, and age and percent gray matter as continuous variables in a general linear model. We tested the full factorial model and also reduced models with the four main effects without any interactions. This allowed us to infer which variables should be retained in the model testing for genotype. When the effect of scanner was significant in the full factorial model or the reduced model with main effects only, the values of the dependent variable were transformed to a Z-score (the mean value for the individual scanner was subtracted from the dependent variable and divided by the standard deviation), thus removing scanner effects before further analyses. The other covariates were then retested with the Z-scores for the dependent variable. If the covariates were not significant in the initial analyses, they were dropped from further consideration. If effects of genotype were significant, the r^2 of the model was used to estimate the proportion of

Table 1 Location and Association with Function of the 10 Genotyped SNPs

SNP[a]	NCBI dbSNP rs no.	Coding strand alleles[b]	CBDB control MAF	Location	Chromosome 2 position[c]	Intermarker distance[d]	Distance from M01	Evidence for				
								Function (mRNA expression)	Clinical association	Association with cognition	Association with gray matter loss	Association with fMRI activation
M01	rs10432420	G/A	0.31	5′ Flanking	171 485 819	0	0		scz (Straub et al, 2007)	(Straub et al, 2007); (Addington et al, 2005)		
M02	rs7557793	T/C	0.27	5′ Flanking	171 490 832	5013	5013			(Straub et al, 2007)	(Addington et al, 2005)	(Straub et al, 2007)
M03	rs1978340	G/A	0.29	5′ Flanking	171 495 628	4796	9809		scz (Straub et al, 2007); bpd (Lundorf et al, 2005); etoh dep (Kuo et al, 2009)			
M04	rs872123	T/C	0.25	5′ Flanking	171 496 673	1045	10854	(Huang et al, 2007b)	bpd (Lundorf et al, 2005)		(Addington et al, 2005)	
M07	rs3762555	G/C	0.25	5′ Flanking	171 497 902	317	12083			(Straub et al, 2007)		
		Start of transcript NM_000817:			171 498 707		12888					
M09	rs3749034	G/A	0.24	Exon 1, 5′ UTR	171 498 982	695	13163	(Straub et al, 2007)	scz (Straub et al, 2007); cos (Addington et al, 2005); etoh dep (Loh el et al, 2006)	(Straub et al, 2007)		
		Start Codon (Met) of transcript NM_000817:			171 500 609–11		14790-92					
M12	rs2241165	A/G	0.25	Intron 2	171 503 886	3683	18067	(Huang et al, 2007b)	cos (Addington et al, 2005); ped bpd (Geller et al, 2008); etoh dep (Kuo et al, 2009)	(Straub et al, 2007)	(Addington et al, 2005)	
M13	rs11542313[e]	T/C	0.41	Exon 3, His37His	171 504 132	246	18313		scz (Straub et al, 2007)	(Straub et al, 2007)		
M16	rs769390	A/C	0.27	Intron 6	171 518 962	4862	33143		scz (Straub et al, 2007); (Loh el et al, 2006)			
M18	rs3791850	G/A	0.22	Intron 12	171 533 607	5620	47788		scz (Straub et al, 2007); cos (Addington et al, 2005); neuroticism (Hettema et al, 2006); etoh dep (Kuo et al, 2009)	(Straub et al, 2007)		

Abbreviations: MAF, minor allele frequency; scz, schizophrenia; cos, childhood onset schizophrenia; bpd, bipolar disorder; ped bpd, pediatric bipolar disorder; etoh dep, alcohol dependence.
[a]Numbering refers to Straub et al (2007).
[b]Major/minor.
[c]From UCSC May 2004.
[d]In base pairs.
[e]This was formerly rs16823181.

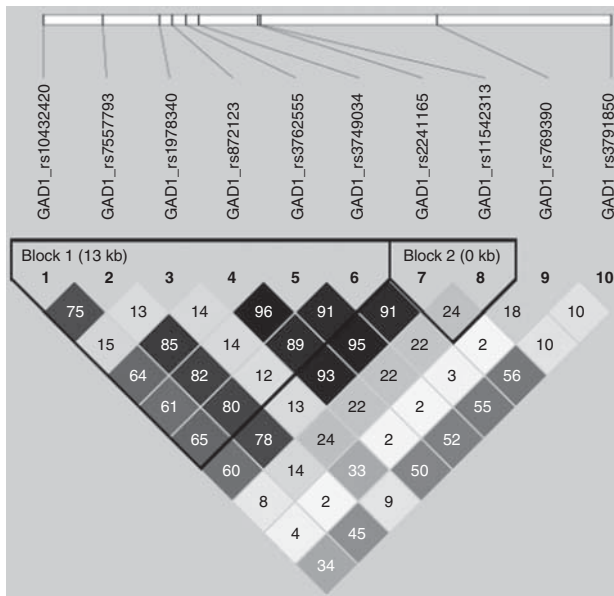**Figure 3** Linkage disequilibrium (r^2) between the 10 genotyped single nucleotide polymorphisms (SNPs). Figure obtained with Haploview (http://www.broadinstitute.org/mpg/haploview).

total variance accounted for by each variable. Separate analyses were repeated for the six *GAD1* SNPs and *COMT* rs4680. In addition, the separate hypothesis was tested that *COMT* rs4680 would interact with rs1978340, rs11542313, and rs769390 in GAD1 based on the clinical results reported by Straub *et al* (2007). For these analyses, to obtain more than two individuals per cell, *GAD1* genotypes were collapsed to two groups: major allele homozygotes and minor allele carriers. As a further assurance against the effects of using different scanners to collect the data, we also tested whether the different genotypes were equally distributed across scanners with a χ^2-statistic.

The *p*-values for the analysis of individual SNPs are reported without correction, but the threshold is $p = 0.007$ for a Bonferroni correction for seven multiple comparisons and $p = 0.017$ for the three multiple comparisons performed to test interactions of GAD1 and COMT val158met. Because each of the hypotheses tested in this report is based on prior evidence of association with clinical diagnosis and on an extensive basic science literature about the roles of GAD67 and dopamine on GABA function, and the current sample is independent of the earlier results, these hypothesis-driven analyses would arguably not be appropriate for agnostic statistical correction.

Because significant effects of GABA/Cre levels could depend on changes in GABA, Cre, or both, we also tested Cre obtained through LCModel in the same manner as described above. A lack of a significant genotype effect for Cre strengthens the case for GABA being the main contributor to the genotype effect. Finally, we tried to rule out other potential covariates such as intelligence quotient (IQ) and years of education by running one-way ANOVAs between genotype groups.

RESULTS

For GABA/Cre, no covariate was significant for the full factorial model, but the model that included only main effects showed significance for age and scanner (Supplementary Table S1). After transforming the GABA/Cre values to *Z*-scores, only age remained significant in the main effects model. Thus, all further analyses of genotype were conducted with z-transformed values and age as a covariate.

Two of the six *GAD1* SNPs showed a significant effect of genotype on GABA/Cre: rs1978340 ($F_{(2,111)} = 5.4$, $p = 0.006$) and rs769390 ($F_{(2,111)} = 5.75$, $p = 0.004$). Both of these associations passed the criteria set by Bonferroni correction for the seven SNP multiple comparisons. *COMT* Val158Met was also nominally significant in this model ($F_{(2,109)} = 3.4$, $p = 0.04$) (but see below for other model results). In all cases, interaction effects with age were not significant and were removed from the model. Supplementary Table S2 shows the complete results. Figure 4 shows the direction of the genotype effects, with the homozygotes for the major allele always to the left of the plot. Age accounted for about 5% and individual *GAD1* SNPs contributed about 8% (Supplementary Table S3) of the total variation in GABA/Cre Z-scores. *COMT* Val158Met accounted for about 7% of the total variance.

Interactions between the three SNPs in *GAD1* (rs1978340, rs11542313, and rs769390) that had shown interaction with *COMT* in the clinical sample were investigated next, after collapsing the genotype groups to have more than two individuals per cell. The interaction effect for the model including rs11542313, *COMT*, and age (main effect of rs11542313: $F_{(1,106)} = 10.6$, $p < 0.0015$ and *COMT*: $F_{(2,106)} = 6.09$, $p < 0.003$) showed a trend toward significance ($F_{(2,106)} = 3.03$, $p = 0.05$: Supplementary Table S2, corrections for multiple comparisons would require at least a *p*-value of 0.017). The effects of genotype are shown in Figure 5. R^2 analysis indicated that the effect of rs11542313 (C allele carriers were combined in one group for this analysis) and of the interaction term each accounted for 4% of the variance in GABA/Cre (Supplementary Table S3). The full model r^2 was 0.21.

Scanner and age were significant covariates in predicting Cre values determined with LCmodel (Supplementary Table S1). After correcting for scanner effects, only age was a significant covariate, thus all significant genotype effects were tested with age as a covariate. No significant main effect of genotype or *GAD1–COMT* interactions was found for Cre Z-scores (Supplementary Table S2). No significant differences for IQ, years of education, or percent gray matter in the spectroscopic voxel were present across *GAD1* or *COMT* genotypes. Finally, genotypes were equally distributed across different scanners ($1.78 < \chi^2 < 5.5$, $0.78 > p > 0.24$ for *GAD1* and *COMT*, respectively).

DISCUSSION

We report a set of positive associations of SNPs in the genes for GAD1 and for COMT and weak evidence of an interaction of these genes on levels of GABA measured in the living human cingulate cortex. These genes are involved in GABA synthesis and GABA neuronal activity, and thus are hypothesis driven and biologically lawful predictors of variation in GABA concentrations. Moreover, the SNPs and the interactions studied here were previously shown to increase risk for the diagnosis of schizophrenia
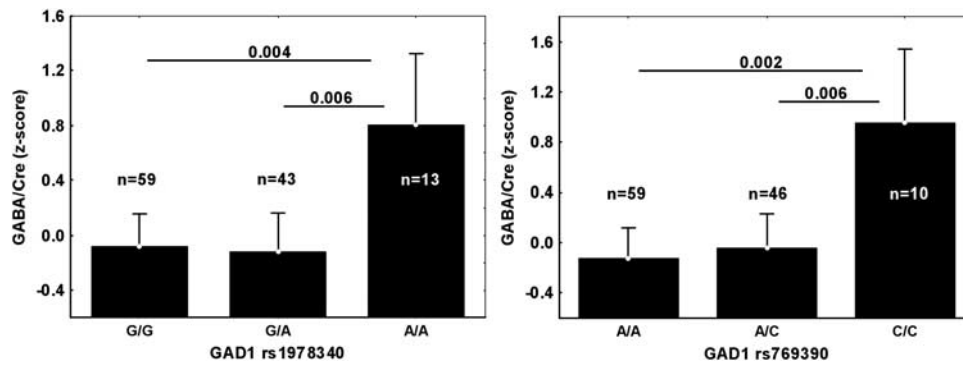
Figure 4 Directionality of main effects of *GAD1* single nucleotide polymorphisms (SNPs) on GABA/Cre. The alleles associated with increased risk for schizophrenia according to Straub *et al* (2007) were A for rs11978340 and C for rs769390. Error bars represent 95% confidence limits. The *p*-values for *post hoc* comparisons (Tukey's 'honest significant difference' correction for multiple comparisons) are also shown, with lines connecting the bars that showed significant differences. All results were adjusted for age in an ANCOVA model.
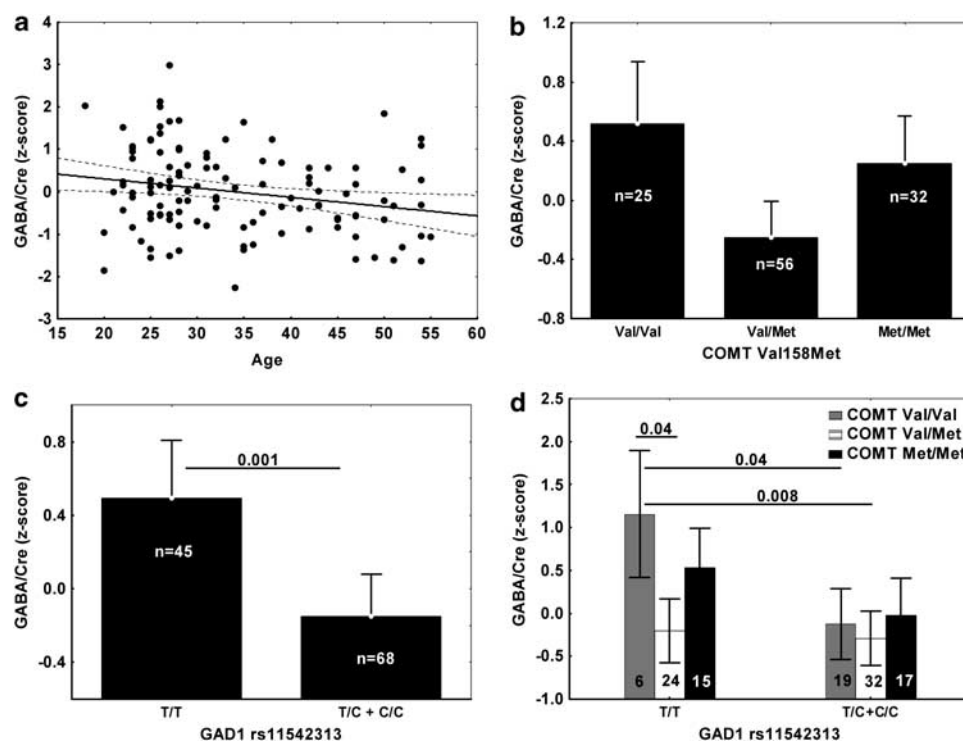

Figure 5 Effects of age (a), *COMT* Val158Met (b), *GAD1* rs11542313 (c), and the interaction of the two genotypes (d). Error bars represent 95% confidence intervals, as does the dotted line in (a). The significant *post hoc* comparisons are also shown, as in Figure 4. The number of individuals in each cell is specified in (c) and (d).

in an independent clinical genetic association data set. Although we have performed a number of statistical tests in this report, and only a few of the results would pass agnostic statistical correction, each test was predicted strongly by the basic science data about GABA and by previous independent clinical association with schizophrenia.

Effects of Genetic Variation in GAD1 on GABA

On the basis of previous evidence of association of SNPs in *GAD1* with schizophrenia and related brain phenotypes, we found significant effects of rs1978340 and rs769390 on cingulate cortex GABA concentrations *in vivo* (Figure 4).

On the basis of *post hoc* tests of significance shown in Figure 4, we found that increased GABA/Cre levels appeared only in homozygotes for the minor allele, suggesting a recessive effect. These results passed the criteria set by Bonferroni correction for the seven SNPs tested. rs11542313 was significant only after collapsing the minor allele carriers, consistent with a recessive effect, and after using a model including *COMT*. These changes were independent of aging (see Figure 5a). In all cases, homozygosity for the *GAD1* alleles thought to confer risk for schizophrenia on the background of *COMT* Val-Val homozygosity in our larger clinical sample (A in rs1978340, C in rs769390, and T for rs11542313; Straub *et al*, 2007) was associated with relatively increased GABA/Cre. The three SNPs identified

as having an effect on GABA measures were in negligible LD ($r^2 = 0.2$–0.3) with each other.

Various mechanisms may account for these effects: changes in the efficiency of production of the protein by alteration of the mRNA secondary structure (Nackley *et al*, 2006) in the case of rs11542313 (a synonymous coding SNP); being part of a regulatory element in the 5′ flanking region in the case of rs1978340 (Veyrieras *et al*, 2008), or in an intron for rs769390; monitoring other rare causative SNPs by LD; or modification of interaction with DNA methyltransferases or with histones thought to regulate epigenetic control of *GAD1* mRNA expression (Kundakovic *et al*, 2007, 2009) and to be relevant for schizophrenia pathology (Huang and Akbarian, 2007a). It should be noted that neither one of these SNPs was associated with alteration in the amount of mRNA for *GAD1* measured in post-mortem brain (Straub *et al*, 2007), however this prior finding was in a sample of mainly African-American ancestry where genetic variation is likely to be associated with different causative haplotypes. Clearly, our study results await replication in an independent sample and a biological explanation for their effect.

For all significant SNPs in *GAD1*, the 'risk' allele was associated with an unexpected increase in GABA/Cre. We would have expected the opposite directionality given the post-mortem evidence for reductions in *GAD1* mRNA expression (see Akbarian and Huang, 2006; Lewis *et al*, 2005 for reviews) and protein levels (Guidotti *et al*, 2000) in schizophrenia (but see Dracheva *et al*, 2004; Gluck *et al*, 2002; Hakak *et al*, 2001, for results in the opposite direction in elderly patients). It is unclear what mechanism might account for this phenomenon. We speculate that GABA production by *GAD1* is actually reduced in risk allele carriers and that a compensatory change such as excessive production by GAD2 or decreased catabolism by GABA transaminase may be causing this unpredicted increase in GABA/Cre levels. The compensation could also be occurring in certain cortical layers, whereas others may be decreased but beyond the spatial resolution of our technique; or be present intracellularly, but correspond to a relative deficit of GABA synaptic transmission; or be limited to calretinin-positive cells that are thought to constitute the majority of GABA-containing cells (Conde *et al*, 1994) and therefore may provide the bulk of the signal for this technique. Alternatively, it is entirely possible that the same genetic variations may show opposite or no effects in patients with schizophrenia. We will be exploring this in future studies. Given the reported association of GAD1 SNPs with neuroticism (Hettema *et al*, 2006), we also checked if our study results may help elucidate the biology of the associations reported by Hettema *et al* (2006), but no significant SNP overlapped between the two studies.

It is also relevant to point out that GAD67 can be regulated by post-transcriptional mechanisms (Martin and Rimvall, 1993; Rimvall and Martin, 1994), and that this may occur in the absence of a change in mRNA levels, a possible mechanisms for a dissociation between GABA and *GAD1* mRNA levels in schizophrenia. If our study results were to be replicated in an independent sample, detailed post-mortem studies of the effects of *GAD1* genetic variation on various markers of GABA interneurons will be necessary to understand the physiology of this phenomenon.

Effects of Genetic Variation in *COMT* on GABA

COMT Val158Met was nominally significant on its own, but became significant beyond our threshold for multiple comparisons when tested in combination with rs11542313, indicating that the two SNPs determine GABA concentrations independently of each other and share separate portions of the variance in GABA values, likely a demonstration of independent biological mechanisms affected by these two genes. *COMT* val158met genotype also showed a trend toward an interaction with *GAD1* rs11542313, but this trend did not survive stringent multiple comparisons correction. Nevertheless, the association of genetic variation in *COMT* and GABA levels and the interaction of *GAD1* and *COMT* genotypes on GABA levels were strikingly consistent with basic science data that cortical DA regulates GABA neuronal activity. The effects of *COMT* rs4680 contributed 6% of the total variation in GABA/Cre, compared with 8% for rs769390 and rs1978340 in *GAD1*, consistent with a modulatory effect. However, the pattern of genotype associations accounting for the main effect of *COMT* was unexpected, with Val/Met heterozygotes showing the lowest GABA/Cre ratios. We had expected Val-Val homozygotes (having the lowest cortical synaptic DA levels) to show the lowest levels of GABA. Although the higher levels of GABA in Met homozygotes might have been expected based on higher dopamine being related to a net increase in inhibitory effect in non-fast spiking interneurons (mainly situated in the superficial layers of the cortex and containing calretinin: Gao *et al*, 2003b), the higher GABA level in Val homozygotes is difficult to explain. The interaction effect between *COMT* and *GAD1* rs11542313 was mainly promoted by the six T/T homozygotes (the *GAD1* 'risk' genotype) who were also Val-Val homozygotes (Figure 5d). Given this limitation, we are inclined to await further replication or a substantially larger sample size before interpreting the finding. Although one possible explanation for these counterintuitive findings may lie in the distribution of D1 and D5 receptor subtypes on different subtypes of GABAergic neurons (Glausier *et al*, 2008), or in a homeostatic mechanism to maintain optimal levels of noise or tuning of the system (Rolls *et al*, 2008), not all aspects of our findings can be explained by these models of DA and GABA receptor modulation of working memory.

Methodological Strengths and Limitations

This study had several methodological strengths. First, our spectroscopic technique allowed for highly accurate separation of the GABA peak from other signals. This same methodology has also been shown to be sensitive to the effects of disease and has shown expected reductions of GABA levels in depression (Hasler *et al*, 2007). Moreover, preliminary data (Geramita *et al*, unpublished observations) show excellent reproducibility in repeated sessions (around 6–8% coefficient of variation) for GABA/Cre. Second, this sample was exclusively composed of Caucasian volunteers, limiting genetic heterogeneity and improving our ability to discern effects due to allelic variation. Finally, special care was put into removing individuals who had any history of psychiatric disorders or substance abuse.

Limitations of this study include that it was conducted on multiple scanners, due to forced changes in platforms during the collection of the sample. It should be kept in mind that the genotypes of the significant SNPs were equally distributed across scanners. Moreover, we cannot account for differences in macromolecules (contributing up to 56% of the GABA peak: Kegeles et al, 2001; van der Veen et al, 2007) between genotypes, but given that the macromolecule signal is formed by many different lipoprotein species, it is highly unlikely that a systematic genetic effect would be found.

Difficulties in interpretation also arise when considering ratios to Cre, however there were no effects of genotype when considering absolute values of Cre obtained by analyzing the unedited portion of the spectrum with LCmodel (Provencher, 2001), supporting the contention that this effect is due mostly to GABA. Conversely, the effect of age on GABA/Cre levels appears to be mainly explained by rising Cre values, a finding consistent with some reports (Angelie et al, 2001; Maudsley et al, 2009; Pfefferbaum et al, 1999). The contribution of partial voluming of gray and white matter in the voxel was also addressed by the use of GABA/Cre ratios as there was no statistical effect of percent gray matter in the voxel in this large sample of controls. Moreover, there were no significant differences in the percent of gray matter present in the voxel between genotypes.

Conclusion

In conclusion, we showed for the first time that genetic variation in *GAD1* and in *COMT* was associated with alterations in GABA/Cre ratios *in vivo*. The directionality of the genetic relationship was opposite to that predicted. It is likely that the association detected here is determined early in development, because none of our analyses showed a significant interaction of age and genotype, therefore GABA/Cre levels are likely to represent a relatively stable trait. Our study results might be the first *in vivo* evidence supporting a role of GAD67 in the development of GABA circuitry (Chattopadhyaya et al, 2007) and subsequent determination of GABA levels.

DISCLOSURE

This work was funded entirely by the NIMH IRP. The authors declare that, except for income received by the primary employer (the NIH), no financial support or compensation has been received from any individual or corporate entity over the past 3 years for research or professional services and there are no personal financial holdings that could be perceived as constituting a potential conflict of interest.

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Supplementary Information accompanies the paper on the Neuropsychopharmacology website (http://www.nature.com/npp)